

Chebyshev

Theorem 1.1 (*Chebyshev*) Suppose X is a random variable with finite expectation, μ , and variance, σ^2 . Then for any real $\epsilon > 0$

$$P(|X - \mu| \geq \epsilon) \leq \frac{\sigma^2}{\epsilon^2} \quad (1)$$

or

$$P(|X - \mu| < \epsilon) \geq 1 - \frac{\sigma^2}{\epsilon^2} \quad (2)$$

The following figure illustrates this point.

Figure 1: Chebyshev

Proof 1.1 *By definition*

$$\sigma^2 = \int_{-\infty}^{\infty} (x - \mu)^2 f_X(x) dx \quad (3)$$

Following Figure 1, we have

$$\sigma^2 = \int_{-\infty}^{\mu-\epsilon} (x-\mu)^2 f_X(x) dx + \int_{\mu-\epsilon}^{\mu+\epsilon} (x-\mu)^2 f_X(x) dx + \int_{\mu+\epsilon}^{\infty} (x-\mu)^2 f_X(x) dx \quad (4)$$

Note that

$$\int_{\mu-\epsilon}^{\mu+\epsilon} (x - \mu)^2 f_X(x) dx \geq 0 \quad (5)$$

so that

$$\sigma^2 \geq \int_{-\infty}^{\mu-\epsilon} (x - \mu)^2 f_X(x) dx + \int_{\mu+\epsilon}^{\infty} (x - \mu)^2 f_X(x) dx \quad (6)$$

As x tends from $-\infty$ to $\mu - \epsilon$ and from $\mu + \epsilon$ to ∞ the smallest values of $(x - \mu)^2$ are attained for $\mu \pm \epsilon$. Hence,

$$\sigma^2 \geq \int_{-\infty}^{\mu-\epsilon} \epsilon^2 f_X(x) dx + \int_{\mu+\epsilon}^{\infty} \epsilon^2 f_X(x) dx \quad (7)$$

$$= \epsilon^2 \left[\int_{-\infty}^{\mu-\epsilon} f_X(x) dx + \int_{\mu+\epsilon}^{\infty} f_X(x) dx \right] \quad (8)$$

$$= \epsilon^2 [P(X \leq \mu - \epsilon) + P(X \geq \mu + \epsilon)] \quad (9)$$

$$= \epsilon^2 P(|X - \mu| \geq \epsilon) \quad (10)$$

Thus

$$P(|X - \mu| \geq \epsilon) \leq \frac{\sigma^2}{\epsilon^2} \quad (11)$$

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Alternatively, letting $\epsilon = h\sigma$ for $h > 0$ we have

$$P(|X - \mu| \geq h\sigma) \leq \frac{1}{h^2} \quad (12)$$

Choosing a different h allows us to find upper bounds on the probability that X takes on a value outside the interval $\mu \pm h\sigma$. We get that, irrespective of the distribution of X , this probability is at most $\frac{1}{h^2}$.